# Differentially Private Hierarchical Count-of-Counts Histograms 

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August 302018

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## Outline

(1) Introduction: hierarchical count-of-counts histograms
(2) Non-hierarchical count-of-counts histograms publishing
(3) Hierarchical count-of-counts histograms publishing
4) Experimental results

## Scenario

- Table Persons(person_name, group_id, location)
- A hierarchy $\Gamma$ on location associated with each group

| name | g_id | loc. |
| :---: | :---: | :---: |
| Alice | 1 | a |
| Bob | 1 | a |
| Carol | 1 | a |
| Dave | 1 | a |
| Eve | 2 | b |
| Frank | 2 | b |
| Judy | 3 | a |
| Nick | 4 | b |

Queries: In the United States,


- How many groups have size 1 ?
- How many groups have size 2 ? In New York,
- How many groups have size 1 ?
- How many groups have size 2 ?

Application:
(1) group $=$ a taxi, data item $=$ a pick up, size $=\#$ of pickup
(2) group $=$ a census block, data item $=$ a person of a specific race, size $=\#$ people of a specific race

## Convenient Views of the Dataset

- A=SELECT groupid, COUNT(*) AS size FROM Persons GROUPBY groupid
- H=SELECT size, COUNT(*) FROM A GROUPBY size

SQL query resulting table A:

| g_id | size | loc. |
| :---: | :---: | :---: |
| 1 | 4 | a |
| 2 | 2 | b |
| 3 | 1 | a |
| 4 | 1 | b |

- count-of-counts histogram (coco) $H$ is

$$
\begin{aligned}
& H^{\text {root }}=[2,1,0,1] \\
& H^{a}=[1,0,0,1]
\end{aligned}
$$

- unattributed histogram [HRMS10] $\mathrm{H}_{\mathrm{g}}$ is

$$
\begin{aligned}
& H_{g}^{\text {root }}=[1,1,2,4] \\
& H_{g}^{a}=[1,4]
\end{aligned}
$$

- cumulative count-of-counts histogram $H_{c}$ is $H_{c}^{\text {root }}=[2,3,3,4]$ $H_{c}^{a}=[1,1,1,2]$


## Protect Privacy

## Definition (Differential Privacy [DMNS06])

A mechanism $M$ satisfies $\epsilon$-differential privacy if, for any pair of databases $D_{1}, D_{2}$ that differ by the presence or absence of one record in the Persons table, and for any possible set $S$ of outputs of $M$, the following is true:

$$
P\left(M\left(D_{1}\right) \in S\right) \leq e^{\epsilon} P\left(M\left(D_{2}\right) \in S\right)
$$

## Geometric Mechanism

## Definition (Sensitivity)

Given a query $q$ (which outputs a vector), the global sensitivity of $q$, denoted by $\Delta(q)$ is defined as:

$$
\Delta(q)=\max _{D_{1}, D_{2}}\left\|q\left(D_{1}\right)-q\left(D_{2}\right)\right\|_{1}
$$

where databases $D_{1}, D_{2}$ contain the public Hierarchy and Groups tables, and differ by the presence or absence of one record in the Persons table.

## Definition (Geometric Mechanism [GRS09])

Given a database $D$, a query $q$ that outputs a vector, a privacy loss budget $\epsilon$, the global sensitivity $\Delta(q)$, the geometric mechanism adds independent noise to each component of $q(D)$ using distribution: $P(X=k)=\frac{1-e^{-\epsilon}}{1+e^{-\epsilon}} e^{-\epsilon|k| / \Delta(q)}$ (for $k=0, \pm 1, \pm 2$, etc.). This distribution is known as the double-geometric with scale $\Delta(q) / \epsilon$.

## Problem Definition

For each node $\tau$ in hierarchy $\Gamma$, create differentially private estimate $\tau . \hat{H}$ of count-of-counts histogram $H$ such that

- $\tau . \hat{H}$ is a count-of-counts histogram (its entries are nonnegative integers)
- The counts are accurate ( $\tau . \widehat{H}$ and $\tau . H$ are close)
- $\tau . \widehat{H}$ matches publicly known total number of groups $G$ in $\tau$
- satisfy consistency: children histograms sum up to the parent.



## Error Measure

- The Earthmover's distance (emd): the minimum number of people that must be added or removed from groups in $\tau . H$ to get $\tau . \widehat{H}$.


$$
\mathrm{emd}=\left|H_{c}-\widehat{H_{c}}\right|_{1}=\left|H_{g}-\widehat{H_{g}}\right|_{1}=2
$$

## Lemma ([NLV07])

The earthmover's distance between $H$ and $\widehat{H}$ can be computed as $\left\|H_{c}-\widehat{H}_{c}\right\|_{1}$, where $H_{c}\left(\right.$ resp., $\left.\widehat{H}_{c}\right)$ is the cumulative histogram of $H$ (resp., $\widehat{H}$ ). It is the same as the $L_{1}$ norm in the $H_{g}$ representation when the number of groups is fixed.

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## Naive Strategy


(1) $\widetilde{H}$ : Add independent double-geometric noise with scale $2 / \epsilon$ to each element of coco histogram $H$
(2) Post-process $\widetilde{H}$ with optimization problem:

$$
\begin{aligned}
& \widehat{H}=\arg \min _{\widehat{H}}\|\widetilde{H}-\widehat{H}\|_{2}^{2} \\
& \text { s.t. } \widehat{H}[i] \geq 0 \text { for all } i \quad \text { and } \sum_{i} \widehat{H}[i]=G
\end{aligned}
$$

(3) To get integers, we set $r=G-\sum_{i}\lfloor\widehat{H}[i]\rfloor$, round the cells with the $r$ largest fractional parts up, and round the rest down.
(9) Solver: quadratic program (e.g., Gurobi [GO16])

## Unattributed Histogram [HRMS10] $\mathrm{H}_{\mathrm{g}}$


(1) Convert coco histogram $H \Rightarrow$ unattributed histogram $H_{g}$
(2) $\widetilde{H}_{g}$ : Add independent double-geometric noise with scale $1 / \epsilon$ to each element of $\mathrm{H}_{g}$
(3) Post-process with optimization problem with either $p=1$ or $p=2$ :

$$
\begin{aligned}
& \widehat{H}_{g}=\arg \min _{\hat{H}_{g}}\left\|\widetilde{H}_{g}-\widehat{H}_{g}\right\|_{p}^{p} \\
& \quad \text { s.t. } 0 \leq \widehat{H}_{g}[i] \leq \widehat{H}_{g}[i+1] \text { for } i=0, \ldots, G-1
\end{aligned}
$$

(9) Round each entry of $\widehat{H}_{g}$ to the nearest integer and convert it back to $\widehat{H}$
(3) Solver: min-max algorithm [BB72], pool-adjacent violators (PAV) [BBBB, RW ${ }^{+} 68$ ], Gurobi [GO16]

## Cumulative Sum Histograms $\mathbf{H}_{\mathbf{c}}$


(1) Convert coco histogram $H \Rightarrow$ cumulative sum histogram $H_{c}$
(2) $\widetilde{H}_{c}$ : Add independent double-geometric noise with scale $1 / \epsilon$ to each element of $H_{c}$
(3) Post-process with optimization problem with either $p=1$ or $p=2$ :

$$
\begin{aligned}
\widehat{H}_{c} & =\arg \min _{\hat{H}_{c}}\left\|\widehat{H}_{c}-\widetilde{H}_{c}\right\|_{p}^{p} \\
& \text { s.t. } 0 \leq \widehat{H}_{c}[i] \leq \widehat{H}_{c}[i+1] \text { for } i=0, \ldots, K \\
& \text { and } \widehat{H}[K]=G
\end{aligned}
$$

(4) Round each entry of $\widehat{H_{c}}$ to the nearest integer and convert it back to $\widehat{H}$
© Solver: min-max algorithm [BB72], pool-adjacent violators (PAV) [BBBB, RW ${ }^{+} 68$ ], Gurobi [GO16]

## Methods Summary

- Naive approach had several orders of magnitude worse error than the unattributed histogram $\mathrm{H}_{\mathrm{g}}$ and cumulative sum histogram $\mathrm{H}_{\mathrm{c}}$ method
- For most datasets, $\mathrm{H}_{\mathrm{c}}$ method generally performs better
- For sparse datasets, $\mathbf{H}_{\mathbf{g}}$ method is better


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## Non-hierarchical Methods Issue

- Estimate coco histograms at each node $\tau, c_{1}, c_{2}$
- Drawback: parent $\tau . \widehat{H}$ does not equal to the sum of children $\left(c_{1} \cdot \widehat{H}+c_{2} \cdot \widehat{H}\right)$


Hierarchical count-of-counts histograms publishing

## Mean-Consistency Algorithm [HRMS10]

(1) Take cumulative coco histograms $H_{c}$ at every node
(2) Add independent double-geometric noise with scale $1 / \epsilon$ to each element of $H_{c}$
(3) Post-process with mean-consistency algorithm


- Drawback: counts can be negative and fractional


## Bottom-up Aggregation

(1) Estimate coco histogram $H$ only at the leaves
(2) Aggregate them up the hierarchy


- Drawback: it introduces high error at non-leaf nodes (like in other hierarchical problems [HRMS10, QYL13])


## Consistency Solution

- Our proposed solution:
(1) Converts estimated coco $\tau . \widehat{H} \Rightarrow$ the unattributed histogram $\tau . \widehat{H}_{g}$
(2) Find a 1-to-1 optimal matching between groups at the child nodes and groups at the parent node
(3) Merge those two estimates


Figure: Before matching


Figure: Consistency result

## Optimal Matching Algorithm

- For each node $\tau$ and its children, we set up a bipartite weighted graph
- There are $\tau . G$ vertices on the top: $(\tau, 1),(\tau, 2), \ldots,(\tau, \tau . G)$. Each vertices on the bottom has the form $(c, j)$, where $c$ is a child of $\tau$ and $j$ is an index into c. $\widehat{H}_{g}$.
- Edge between every vertex $(\tau, i)$ and $(c, j)$ has weight $\left|\tau . \widehat{H}_{g}[i]-c . \widehat{H}_{g}[j]\right|:$ measure the difference in estimated size

- Our desired matching is least cost weighted matching on this bipartite graph.
- Optimal algorithm: matching the smallest unmatched group in $\tau$ to the smallest unmatched group among any of its children.


## Top-down Consistency



Figure: Level 0 and Level 1 consistency matching

| (1) | (1) | (1) | (1) | (2) | (2) | (3) | (8) | (8) |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |




Figure: Level 1 and Level 2 consistency matching
(1) Consistency matching at top level
(2) Use new estimates for next level consistency
(3) Use the new merged estimates at the leaves for back substitution to get unattributed histogram:

$$
\begin{aligned}
\widehat{H}_{g}^{a} & =[1,1,1,2,9] \\
\widehat{H}_{g}^{b} & =[1,3,3,6] \\
\widehat{H}_{g}^{\text {root }} & =[1,1,1,1,2,3,3,6,9]
\end{aligned}
$$

(3) Convert consist unattributed histogram ing count-of-counts histogram

## Initial Variance Estimation

Recall: we convert $\tau . \widehat{H}$ into the unattributed histogram $\tau . \widehat{H}_{g}$.
For each $i$, we need an estimate of the variance of the $i^{\text {th }}$ largest group $\tau . \widehat{H}_{g}[i]$, so that it can be used to merge two estimates during matching.

- Let $S_{i}$ be the number of groups that were in the same partition as $i$ in the solution

- Let $\epsilon$ be the privacy budget used in node $\tau$ in level $\ell$ of $\Gamma$

For the $\mathbf{H}_{\mathbf{g}}$ method:

- Variance estimate for the $i^{\text {th }}$ largest group: $\tau . V_{g}[i]=\frac{2}{\left|S_{i}\right| \epsilon^{2}}$ For the $\mathbf{H}_{\mathbf{c}}$ method:
- Variance estimate of the $i^{\text {th }}$ largest group: $\tau . V_{g}[i]=4 /\left(\epsilon^{2} \times\right.$ number of estimated groups of size $\left.\tau . \widehat{H}_{g}[i]\right)$


## Merge Estimates

Given a node $\tau$, the matching algorithm assigns one group $i$ in $\tau$ to one group $j$ in some child of $\tau$
$\Rightarrow$ for every group, two estimates of its size: $\tau . \widehat{H}_{g}[i]$ and $c . \widehat{H}_{g}[j]$ \& estimates of variance $\tau . V_{g}[i]$ and $c . V_{g}[j]$.

- Optimal linear combination of the estimates [HRMS10]: weighted average

$$
\begin{equation*}
\left(\frac{\tau \cdot \widehat{H}_{g}[i]}{\tau \cdot V_{g}[i]}+\frac{c \cdot \widehat{H}_{g}[j]}{c \cdot V_{g}[j]}\right) /\left(\frac{1}{\tau \cdot V_{g}[i]}+\frac{1}{c \cdot V_{g}[j]}\right) \tag{1}
\end{equation*}
$$

and the variance of this estimator is

$$
\begin{equation*}
\left(\frac{1}{\tau . V_{g}[i]}+\frac{1}{c . V_{g}[j]}\right)^{-1} \tag{2}
\end{equation*}
$$

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## Experiments

Use 4 datasets:

- Race distribution - White (2010 Census data [Bur12]): For West Coast/State/County and a given race, for each $j$, how many Census blocks contain $j$ people of that race?
- Race distribution - Hawaiian [Bur12]
- Partially synthetic housing: The number of individuals in each facility is important but this information was truncated past households of size 7 in the 2010 Decennial Census Summary File 1 [Bur12]. We add a heavy tail as would be expected from group quarters (e.g., dormitories, barracks, correctional facilities).
- NYC taxi: In 2013, how many taxis had $j$ pickups in Manhattan/Town/Neighborhood?


## Ruling out Naive Strategy

- Naive strategy's average error is in the billions

Table: Average error with $\epsilon=1.0$ at top level

| Method | Synthetic | White | Hawaiian | Taxi |
| :---: | :---: | :---: | :---: | :---: |
| Naive | $4,462,728,374$ | $4,809,679,734$ | $4,027,891,692$ | $208,977,518$ |
| $\mathbf{H}_{\mathbf{c}}$ | $3,742.0$ | $1,838.9$ | 254.0 | $2,819.8$ |
| $\mathbf{H}_{\mathbf{g}}$ | $2,219.6$ | $6,115.3$ | 516.2 | $11,227.6$ |

## Weighted average estimation comparison

- Two choices at each level: $\mathbf{H}_{\mathbf{c}}, \mathbf{H}_{\mathbf{g}}$
- Weighted average method consistently produces large reductions in error at the top level


(a) partially synthetic housing (national)

(d) partially synthetic housing (state)

(e) white (state)

(c) hawaiian (national)

(f) hawaiian (state)

Figure: Merging estimates using weighted average vs. normal average. $x$-axis: privacy budget per level.

## Comparison to Bottom-up Aggregation

- Allocate all privacy budget (total privacy budget of $\epsilon=1.0$ in the table) to the leaves and set the coco histogram of a parent to be the sum of the histograms at the leaves.
- Very low error at the leaves but higher error everywhere else

|  | Part. Synth. | White | Hawaiian | Taxi |
| :---: | :---: | :---: | :---: | :---: |
| Level 0 |  |  |  |  |
| BU | 78, 459.0 | 448, 909.0 | 13, 968.0 | 20,731.0 |
| $H_{c}$ | 32,480.0 | 17,000.0 | 1,381.0 | 10,547.0 |
| Level 1 |  |  |  |  |
| BU | 1,512.2 | 8,722.0 | 270.1 | 10, 405.5 |
| $H_{c}$ | 1,000.3 | 1,511.8 | 117.7 | 5,431.5 |
| Level 2 |  |  |  |  |
| BU | 24.9 | 152.3 | 4.3 | 772.8 |
| $H_{c}$ | 80.1 | 363.8 | 21.6 | 1,601.8 |

## 3-Level Hierarchy Results

- Two alternatives $H_{g} \times H_{g} \times H_{g}$ and $H_{c} \times H_{c} \times H_{c}$
- Data dependent performance: $\mathbf{H}_{\mathbf{c}}$ performs better in dense region while


(b) partially synthetic housing (Level 1)

(a) partially synthetic housing (Level 0)



(g) hawaiian (Level 0)

(h) hawaiian (Level 1)



(k) taxi (Level 1)

(1) taxi (Level 2)


## Summary

- Introduced hierarchical count-of-counts problem, along with appropriate error metrics
- Proposed a differentially private solution that generates non-hierarchical and hierarchical version of count-of-counts histograms.
- $\mathbf{H}_{\mathbf{c}}$ method generally performs better on dense dataset while datasets with more sparsity favor $\mathbf{H}_{g}$ method

Questions?

## References I

圊 RE Barlow and HD Brunk．
The isotonic regression problem and its dual．
Journal of the American Statistical Association，pages 140－147， 1972.
R RE Barlow，DJ Bartholomew，JM Bremner，and HD Brunk．
Statistical inference under order restrictions． 1972.
回 U．S．Census Bureau．
2010 census summary file 1,2010 census of population and housing， technical documentation．
https：／／www．census．gov／prod／cen2010／doc／sf1．pdf， 2012.
嗇 Cynthia Dwork，Frank McSherry，Kobbi Nissim，and Adam Smith．
Calibrating noise to sensitivity in private data analysis．
In Proceedings of the Third Conference on Theory of Cryptography，
pages 265－284， 2006.

## References II

國 Inc．Gurobi Optimization．
Gurobi optimizer reference manual， 2016.
围 Arpita Ghosh，Tim Roughgarden，and Mukund Sundararajan．
Universally utility－maximizing privacy mechanisms．
In STOC，pages 1673－1693， 2009.
Richael Hay，Vibhor Rastogi，Gerome Miklau，and Dan Suciu． Boosting the accuracy of differentially private histograms through consistency．
PVLDB，3（1－2）：1021－1032， 2010.
求 T．Li N．Li and S．Venkatasubramanian．
t－closeness：Privacy beyond k－anonymity and I－diversity． In ICDE，pages 106－115， 2007.

## References III

回 Wahbeh Qardaji, Weining Yang, and Ninghui Li.
Understanding hierarchical methods for differentially private histograms.
PVLDB, 6(14):1954-1965, 2013.
圊 Tim Robertson, Paul Waltman, et al.
On estimating monotone parameters.
The Annals of Mathematical Statistics, pages 1030-1039, 1968.

